

# Uncertainty quantification for CH<sub>4</sub> emission estimates in TM5-4DVAR

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TM5-4D-Var is a *deterministic* inversion framework.

- We estimate the  $\text{CH}_4$  emissions that *best fit* the model and observations under some error assumptions (an optimal solution, also called the *analysis*).

However, error covariance information for the optimal solution is not easy to obtain.

- If the model is linear, and (model and observation) errors are Gaussian, then TM5-CONGRAD can estimate the leading components of the analysis error covariance matrix.
- **What can we do if the model is nonlinear and/or prior emission statistics are non-Gaussian?**

We are given:

- the background state vector  $\mathbf{x}_b$  with error covariance  $\mathbf{B} = \mathbf{D}\mathbf{C}\mathbf{D} := \mathbf{L}\mathbf{L}^T$  ( $\mathbf{L} := \mathbf{D}\mathbf{C}^{1/2}$ ),
- the (potentially nonlinear) model operator  $\mathbf{H}$  (forward model + observation operator), and
- the measurement data vector  $\mathbf{y}$  with error covariance  $\mathbf{R}$ .

Solve for the *analysis* state:

$$\mathbf{x}_a = \arg \min_{\mathbf{x}} \mathcal{J} = \frac{1}{2}(\mathbf{x} - \mathbf{x}_b)^T \mathbf{B}^{-1}(\mathbf{x} - \mathbf{x}_b) + \frac{1}{2}(\mathbf{H}\mathbf{x} - \mathbf{y})^T \mathbf{R}^{-1}(\mathbf{H}\mathbf{x} - \mathbf{y})$$

# CONGRAD UQ (Meirink et al., ACP, 2008)

## CONGRAD

Assumptions:  $\mathbf{H}$  is **linear**. The prior and observation errors are **Gaussian**. Then:

$$\nabla_{\mathbf{x}}^2 \mathcal{J} = \mathbf{B}^{-1} + \mathbf{H}^T \mathbf{R}^{-1} \mathbf{H}$$

We optimize the preconditioned state vector  $\boldsymbol{\xi} := \mathbf{L}^{-1} \mathbf{x}$ . The **covariance matrix of the analysis errors** reads:

$$\mathcal{P}^{\text{apos}}(\mathbf{x}_a) = (\nabla^2 \mathcal{J}(\mathbf{x}_a))^{-1} \approx \mathbf{B} + \sum_{i=1}^K \left( \frac{1}{\lambda_i} - 1 \right) (\mathbf{L} \mathbf{v}_i) (\mathbf{L} \mathbf{v}_i)^T$$

$(\lambda_i, \mathbf{v}_i)_{i=1 \dots K}$  are the **leading eigen-pairs** of the preconditioned Hessian matrix.

The **uncertainty reduction factor**:  $1 - \frac{\sigma_{\text{apos}}}{\sigma_{\text{apri}}}$ , with  $\sigma_{\text{apos}} = \sqrt{\text{diag}(\mathcal{P}^{\text{apos}}(\mathbf{x}_a))}$ .

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- ☹ We may get **negative emissions** in regions that are not well constrained by observations (we can fit anything with a Gaussian...).

# TM5-4DVAR: (semi-)lognormal emissions prior

We optimize the *emission factors*  $\mathbf{x}$ , not the absolute emissions  $\mathbf{z}$ :

Semi-lognormal emissions, e.g. Bergamaschi et. al, 2009

$$\mathbf{z}_{\text{emis}}^i := \begin{cases} E_{\text{apri}}^i (1 + \mathbf{x}_{\text{emis}}^i) & \mathbf{x}_{\text{emis}}^i \geq 0 \\ E_{\text{apri}}^i \exp \mathbf{x}_{\text{emis}}^i & \mathbf{x}_{\text{emis}}^i < 0. \end{cases}$$

If  $\mathbf{x} \sim \mathcal{N}(0, \mathbf{B})$ , then  $\mathcal{E}\{\mathbf{z}_{\text{emis}}^i\} = E_{\text{apri}}^i \left( \frac{1}{2} + \frac{1}{\sqrt{2\pi}} \sigma_i + \frac{1}{2} \exp \frac{\sigma_i^2}{2} \operatorname{erfc} \frac{\sigma_i}{\sqrt{2}} \right)$ .

Lognormal emissions

$$\mathbf{z}_{\text{emis}}^i := E_{\text{apri}}^i \exp \mathbf{x}_{\text{emis}}^i$$

If  $\mathbf{x} \sim \mathcal{N}(0, \mathbf{B})$ , then  $\mathbf{z} \sim \log \mathcal{N}(E_{\text{apri}}^i, \mathbf{B})$  and  $\mathcal{E}\{\mathbf{z}_{\text{emis}}^i\} = E_{\text{apri}}^i \exp \frac{\sigma_i^2}{2}$ .

The diagonal elements of  $\mathbf{B}$  are now the *relative* per-category emission uncertainties (variances), as defined in the TM5 rc file.

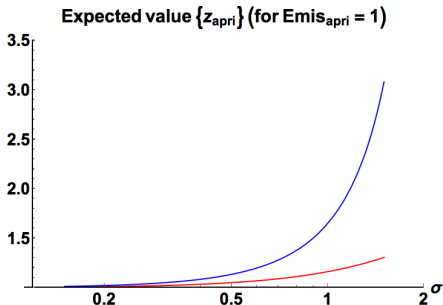
**NB:**  $\mathcal{J}$  is defined in terms of  $\mathbf{x}$ , remains Gaussian (quadratic in  $\mathbf{x}$ ).

# TM5-4DVAR: (semi-)lognormal emissions prior

- ☺ The emission uncertainty space has positive support - a posteriori emissions cannot be negative.

- ☹ We invert for the mean of the (lognormal) a posteriori PDF. But this mean is **unbounded** wrt the (per-category) emission variance. Our choice of PDF pushes us away from the prior  $\mathbf{z}_{\text{apri}}$ .

## Semi-lognormal vs. lognormal



- ☹ The cost function  $\mathcal{J}$  is no longer quadratic in the *absolute* emissions  $\mathbf{z}$ . We have to use M1QN3 instead of CONGRAD.

# UQ with a nonlinear model: M1QN3 Hessian approximation

## Main idea

- What we want: the **leading eigen-pairs**  $(\lambda_i, \mathbf{v}_i)$  of the Hessian matrix  $\nabla_{\xi}^2 \mathcal{J}(\xi_a)$ .
- We cannot form the Hessian matrix explicitly (it can have more than  $\sim 10^{12}$  elements for a zoomed model run). But we don't need to!
- Let's use a *black-box matrix-free eigenvalue solver* (such as ARPACK, available in `scipy`). We just need to calculate the action of the (approximate) Hessian on a given vector  $\mathbf{q}$ .
- TM5-4DVAR implementation: a post-inversion step (after we have computed  $\mathbf{x}_a$ ).
- **How can we approximate the Hessian?**



# The M1QN3 Hessian approximation

## M1QN3

- A limited-memory quasi-Newton optimization routine.
- At iteration  $k$ , the algorithm uses the state and gradient vectors from iterations  $k - 1, \dots, k - m$  ( $m \sim 10$ ) to build an approximation to  $\nabla^2 J(\boldsymbol{\xi})$ .
- ☺ We can efficiently compute the product  $\nabla_{\text{M1QN3}}^2 J(\boldsymbol{\xi}) \times \mathbf{q}$  ( $\mathcal{O}(mn)$  FLOPs).

# Finite-difference (FD) Hessian-vector products

Finite difference of adjoints:  $\epsilon > 0$  (sufficiently small)

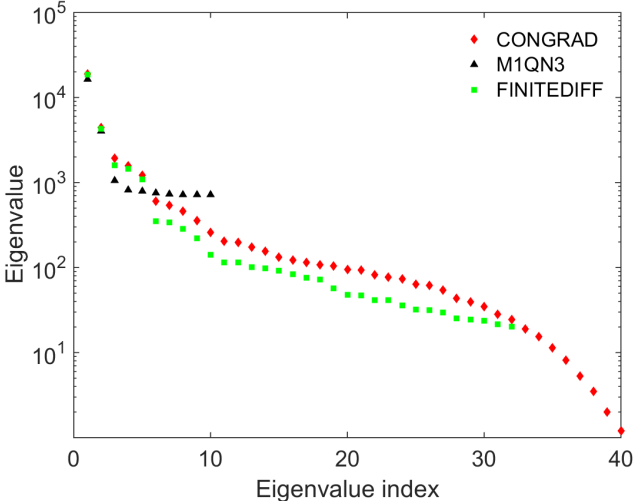
$$\begin{aligned}\boldsymbol{\xi}_a &= \mathbf{L}^{-1}\mathbf{x}_a \\ \nabla_{\boldsymbol{\xi}}^2 \mathcal{J}(\boldsymbol{\xi}_a) \times \mathbf{q} &= \frac{1}{2\epsilon} [\nabla_{\boldsymbol{\xi}} \mathcal{J}(\boldsymbol{\xi}_a + \epsilon \mathbf{q}) - \nabla_{\boldsymbol{\xi}} \mathcal{J}(\boldsymbol{\xi}_a - \epsilon \mathbf{q})] + \mathcal{O}(\epsilon^2)\end{aligned}$$

- ☹ Cost: two adjoint integrations per Hessian-vector product.
- ☺ The adjoint integrations can be performed in parallel.

The analysis error covariance

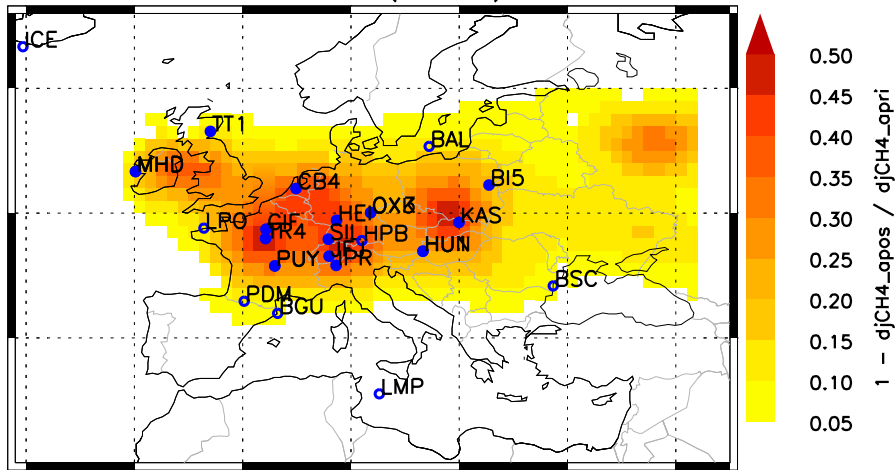
$$\begin{aligned}\mathbf{z}_a &:= h(\mathbf{x}_a) \\ \mathcal{P}^{\text{apos}}(\mathbf{z}_a) &\approx \mathbf{B}_{\text{abs}} + \left( \frac{dh}{d\mathbf{x}} \right) \left( \sum_{i=1}^K \left( \frac{1}{\lambda_i} - 1 \right) (\mathbf{L}\mathbf{v}_i)(\mathbf{L}\mathbf{v}_i)^T \right) \left( \frac{dh}{d\mathbf{x}} \right)^T\end{aligned}$$

# Hessian-based UQ: largest eigenvalues



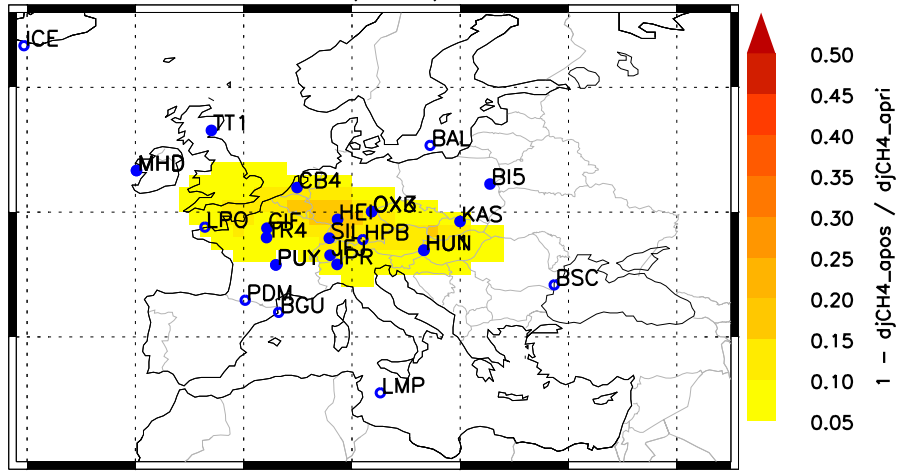
# Hessian-based UQ: uncertainty reduction

CONGRAD relative UR (K=30)



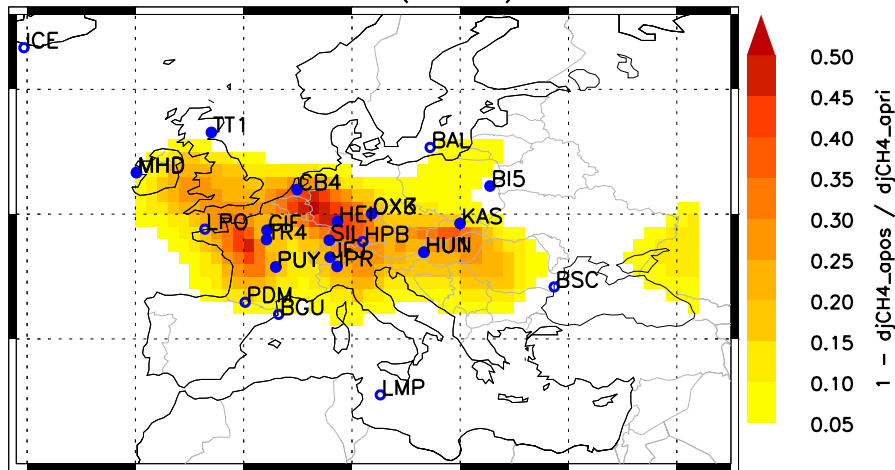
# Hessian-based UQ: uncertainty reduction

M1QN3 relative UR (K=8)



# Hessian-based UQ: uncertainty reduction

FINITEDIFF relative UR (K=30)



# Monte Carlo 4DVAR

## The algorithm

- 1 Generate an ensemble of prior states  $\mathbf{x}_b^{(k)} \sim \mathcal{N}(0, \mathbf{B})$ ,  $k = 1 \dots K$ .
- 2 Perturb point observations and satellite total columns:

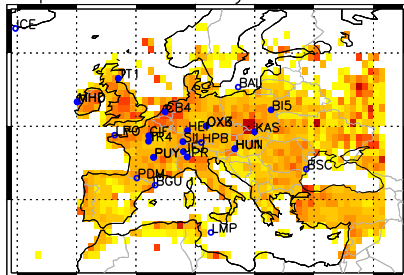
$$\mathbf{y}^{(k)} = \mathbf{y} + \mathbf{R}^{1/2} \boldsymbol{\rho}^{(k)}, \quad \boldsymbol{\rho}^{(k)} \sim \mathcal{N}(0, \mathbf{I}).$$

- 3 Run  $K$  inversions: each starts from  $\mathbf{x}_b^{(k)}$  and assimilates the data vector  $\mathbf{y}^{(k)}$ . We get the a posteriori ensemble  $\mathbf{x}_a^{(k)}$ .
- 4 Calculate a posteriori statistics (mean, variances) from  $\mathbf{x}_a^{(k)}$ , aggregate (absolute) errors with KPP.

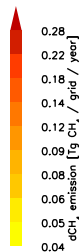
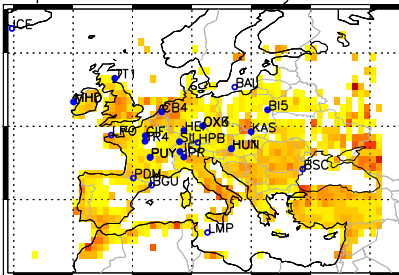
- ☺ Works with any prior/observation PDF, easy to implement (at least in theory), embarrassingly parallel.
- ☹ Computationally expensive, converges slowly (rate  $\sim \frac{1}{\sqrt{K}}$ ).

# Monte-Carlo UQ: numerical results ( $K = 26$ )

a priori uncertainty



a posteriori uncertainty





Computational cost:

**Monte Carlo**

**M1QN3**

**Finite differences**

$$K \times \text{Cost}_{4\text{DVAR}}$$

very small overhead

$$\sim 3 \times K \times [\text{Cost}_{\text{FWD}} + \text{Cost}_{\text{ADJ}}]$$

# Wrap-up

## Computational cost:

<b>Monte Carlo</b>	$K \times \text{Cost}_{4\text{DVAR}}$
<b>M1QN3</b>	very small overhead
<b>Finite differences</b>	$\sim 3 \times K \times [\text{Cost}_{\text{FWD}} + \text{Cost}_{\text{ADJ}}]$

## Code status and future work:

- Monte-Carlo and Hessian-based UQ code now in the TM5-JRC SVN tree, to be merged into the SF .hg branch.
- Calculate and compare uncertainty estimates for a multi-year inversion.
- Make Monte-Carlo more HPC-friendly (use MPI with  $K$  processes instead of  $K$  single-core inversion jobs).
- Invert for the median/mode of the lognormal PDF (previous work by Fletcher and Zupanski, not trivial).

# Questions?



# The nitty-gritty...

## The effect of variable transformations

We are optimizing the emission coefficients  $\mathbf{x}$ .

What we really want is (the eigen-pairs of) the Hessian wrt the absolute emission variables  $\mathbf{z}_{\text{emis}} = h(\mathbf{x})$ . After some math...

$$\frac{\partial^2 \mathcal{J}}{\partial \mathbf{z}^2} \mathbf{q} = \left( \frac{dh}{d\mathbf{x}} \right)^{-T} \frac{\partial^2 \mathcal{J}^2}{\partial \mathbf{x}^2} \left( \frac{dh}{d\mathbf{x}} \right)^{-1} \mathbf{q} - \left( \frac{dh}{d\mathbf{x}} \right)^{-T} \left( \frac{d^2 h}{d\mathbf{x}^2} \otimes \left( \frac{dh}{d\mathbf{x}} \right)^{-1} \mathbf{q} \right)^T \left( \frac{\partial \mathcal{J}}{\partial \mathbf{z}} \right)^T$$

Ignoring the second order term (Hessian of  $h$ ), we get:

$$\mathcal{P}^{\text{apos}}(\mathbf{z}_a) \approx (\nabla^2 \mathcal{J}(\mathbf{z}_a))^{-1} \approx \mathbf{B}_{\text{abs}} + \left( \frac{dh}{d\mathbf{x}} \right) \left( \sum_{i=1}^K \left( \frac{1}{\lambda_i} - 1 \right) (\mathbf{L}\mathbf{v}_i)(\mathbf{L}\mathbf{v}_i)^T \right) \left( \frac{dh}{d\mathbf{x}} \right)^T$$

# Test inversion details

- TM5 version: JRC-M05
- Time period: 12/2009 - 3/2010
- 40 M1QN3 iterations, 40 CONGRAD iterations
- Prior emission uncertainties: 100% for all categories ( $\sigma_x = 1$ ).
- We assimilated surface observations (NOAA) and GOSAT OCPR v52 retrievals.
- Uncertainties aggregated with VPP.